DDPM score matching and distribution learning

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1



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Diffusion models

- Data distribution: $p_{\text{data}} = p_0$.
- Forward process: $dX_t = -X_t dt + \sqrt{2} dB_t$, $X_t \sim p_t$.
- Reverse process: $dX_t^{\leftarrow} = \{X_t^{\leftarrow} + 2\nabla \log p_{T-t}(X_t^{\leftarrow})\} dt + \sqrt{2} dB_t.$

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Training

• Learn score functions via the score matching objective: $\nabla \log p_t = \arg \min_{s_t: \mathbb{R}^d \to \mathbb{R}^d} \mathbb{E}_{x_0 \sim p_0} SM_t(s_t, x_0)$, where

$$SM_t(s_t, x_0)$$

$$\coloneqq \mathbb{E}_{x_t \sim q_{t\mid 0}^{OU}(\cdot \mid x_0)} [\|s_t(x_t)\|^2 + 2\langle s_t(x_t), \nabla \log q_{t\mid 0}^{OU}(x_t \mid x_0) \rangle].$$

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• Empirical version: $\hat{s}_t = \arg\min_{s_t \in \mathcal{S}_t} n^{-1} \sum_{i=1}^n SM_t(s_t, x_0^{(i)}).$

Generation

• Once we have estimated scores $\{\widehat{s}_t\}_{t\in[0,T]}$, we discretize the reverse process: $d\widehat{X}_t^{\leftarrow} = \{\widehat{X}_t^{\leftarrow} + 2\widehat{s}_{t_-}(\widehat{X}_{t_-}^{\leftarrow})\} dt + \sqrt{2} dB_t$, where $\widehat{X}_0^{\leftarrow} \sim \widehat{p}_T = N(0, I_d)$ and $\widehat{X}_T^{\leftarrow} \sim \widehat{p}_0$.

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Theorem: Denoising diffusion models (DDPM) achieve $TV(\hat{p}_0, p_0) \le \varepsilon_*$ in poly $(d, L, 1/\varepsilon_*)$ steps, where:

• $\|\nabla \log p_t\|_{\operatorname{Lip}} \leq L$ and $\mathbb{E}_{x_0 \sim p_0}[\|x_0\|^2] \leq \operatorname{poly}(d)$.

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- Sitan Chen, **S.C**., Jerry Li, Yuanzhi Li, Adil Salim, Anru Zhang, *Sampling is as easy as learning the score*. ICLR 2023.
- [Concurrent] Holden Lee, Jianfeng Lu, Yixin Tan, *Convergence of score-based generative modeling for general data distributions*. ALT 2023.
- Many, many follow-up works...

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- Given samples from p₀, "learn a sampler" for p₀. √ (yes, we shall see that this makes sense)



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- (Parameter Recovery) If 𝒫 = {p_θ : θ ∈ Θ}, the goal is to output an estimate θ̂ of the parameter.
- (Density Estimation) The goal is to output an *evaluator*, i.e., a function p
 : ℝ^d → ℝ such that p
 (x) is a good estimator of the density p(x) at x.
- (Learning a Sampler) The goal is to output a generator, i.e., a function \$\hat{\varsigma}\$: [0, 1] → ℝ^d which takes a random seed U ~ uniform([0, 1]), such that law(\$\hat{\varsigma}\$(U) | X) ≈ p.

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- E.g., minimax lower bounds in statistics apply to both models.
- This is not true for computational lower bounds.

Reinterpreting the DDPM result as distribution learning

Theorem (informal): Let \mathscr{P} be nearly any ("realistic") family of distributions. Then, the sample complexity of learning a sampler for \mathscr{P} is at most the sample complexity of learning the score functions along DDPM for \mathscr{P} .

Moreover, the computational complexity is at most a polynomial factor worse.

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Moreover, the computational complexity is at most a polynomial factor worse.

Learning a sampler is as easy as learning the scores.

Examples of learning a sampler via DDPM

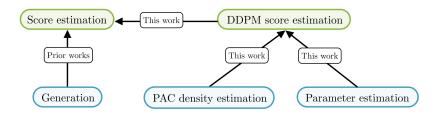
This is a growing literature!

- [Oko, Akiyama, Suzuki '23; Dou, Kotekal, Xu, Zhou '24] Score matching along DDPM yields minimax optimal samplers for Besov and Hölder classes of densities.
- [Chen, Kontonis, Shah '24; Gatmiry, Kelner, Lee '24] Score matching along DDPM yields new algorithmic results for learning mixtures of Gaussians (in the sense of learning a sampler).

• ...

Our new results

[S.C., Alkis Kalavasis, Anay Mehrotra, Omar Montasser, DDPM score matching and distribution learning. '25]



Outline

- A key identity for the likelihood
- Implications for parameter estimation
- Implications for density estimation
- Implications for computational lower bounds

Likelihood identity

Recall:

- $\nabla \log p_t = \arg \min_{s_t: \mathbb{R}^d \to \mathbb{R}^d} \mathbb{E}_{x_0 \sim p_0} \operatorname{SM}_t(s_t, x_0).$
- $\widehat{s}_t = \arg\min_{s_t \in \mathcal{S}_t} n^{-1} \sum_{i=1}^n SM_t(s_t, x_0^{(i)}).$

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Lemma:

$$-\log p_0(x_0) = \int_0^T SM_t(\nabla \log p_t, x_0) dt + C_{d,T} + O(e^{-2T}),$$

where $C_{d,T} = \frac{d}{2} \log(2\pi e (1 - e^{-2T})).$

We do not claim novelty: see, e.g., [Song, Durkan, Murray, Ermon '21; Chen, Liu, Theodorou '22; Li, Yan '24; ...].

Parameter estimation

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Setting: $\mathscr{P} = \{p_{\theta} : \theta \in \Theta\}, \text{ data } x_0^{(1)}, \dots, x_0^{(n)} \stackrel{\text{i.i.d.}}{\sim} p_{\theta^{\star}}.$

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Prior works

• [Koehler, Heckett, Risteski '23] studied the implicit score matching (ISM) estimator: $\widehat{\theta}_n^{\text{ISM}} \coloneqq \arg\min_{\theta \in \Theta} \frac{1}{n} \sum_{i=1}^n \{ \|\nabla \log p_\theta(x_0^{(i)})\|^2 + 2\Delta \log p_\theta(x_0^{(i)}) \}.$ They showed that $\sqrt{n} (\widehat{\theta}_n^{\text{ISM}} - \theta^{\star}) \xrightarrow{d} N(0, \Sigma^{\text{ISM}}).$

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≥ When \mathscr{P} satisfies a "restricted Poincaré inequality", Σ^{ISM} can be bounded in terms of Σ^{MLE} = $\mathcal{I}(\theta^*)^{-1}$.

 \triangleright In general, Σ^{ISM} ≫ Σ^{MLE} (provably inefficient!).

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rapid mixing 🛶 statistical efficiency

• For Gaussian mixtures, [Shah, Chen, Klivans '23; Qin, Risteski '24] established polynomial sample complexity via score matching along other diffusions.

DDPM score matching and parameter estimation

To estimate θ^{\star} , let us minimize the DDPM score matching loss:

$$\widehat{\theta}_n^{\text{DDPM}} \coloneqq \argmin_{\theta \in \Theta} \frac{1}{n} \sum_{i=1}^n \int_0^T SM_t(\nabla \log p_{\theta,t}, x_0^{(i)}) \, \mathrm{d}t$$

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By the likelihood identity:

$$\begin{aligned} \arg\min_{\theta\in\Theta} \left\{ -\frac{1}{n} \sum_{i=1}^{n} \log p_{\theta}(x_0^{(i)}) \right\} \\ &= \arg\min_{\theta\in\Theta} \left\{ \frac{1}{n} \sum_{i=1}^{n} \int_0^T \mathrm{SM}_t(\nabla \log p_{\theta,t}, x_0^{(i)}) \, \mathrm{d}t + C_{d,T} + O(e^{-2T}) \right\} \end{aligned}$$



DDPM score matching achieves full efficiency

Theorem [CKMM '25]: Under standard conditions, provided that $T = T_n$ satisfies $T_n - \frac{1}{2} \log n \rightarrow \infty$,

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Parameter estimation is as easy as (properly) learning scores.

Density estimation

DDPM score matching and density estimation

Likelihood identity:

$$-\log p_0(x_0) = \int_0^T SM_t(\nabla \log p_t, x_0) \, \mathrm{d}t + C_{d,T} + O(e^{-2T}) \, .$$

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An obvious idea is to estimate $-\log p_0(x_0)$ by outputting

$$-\log \widehat{p}_0(x_0) \coloneqq \int_0^T \mathsf{SM}_t(\widehat{s}_t, x_0) \, \mathrm{d}t + C_{d,T} \, .$$





- $\|\nabla \log p_0\|_{\operatorname{Lip}} \leq L$ and $\mathbb{E}_{x_0 \sim p_0}[\|x_0\|^2] \leq \operatorname{poly}(d)$.
- $\int_0^T \|\widehat{s}_t \nabla \log p_t\|_{L^2(p_t)}^2 dt \le \widetilde{O}(\varepsilon^2/d).$

Theorem [CKMM '25]: The DDPM density estimator achieves $\mathbb{E}_{x_0 \sim p_0} |\log(\widehat{p}(x_0)/p(x_0))| \le \varepsilon$ in poly $(d, L, 1/\varepsilon)$ time, where:

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PAC density estimation is as easy as learning the scores.

Example: application to the Hölder class

Let $\mathscr{H}_s(C, L)$ consist of Hölder densities on [-1, 1]. (Here, s = smoothness, C = lower bd. on density, L = size of Hölder ball.)

[Dou, Kotekal, Xu, Zhou '24] obtained optimal rates of score estimation for $\mathscr{H}_s(C, L)$, leading to a minimax optimal sampler.

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Theorem [CKMM '25]: There is a density estimator based on DDPM score matching such that for the L^1 risk $\mathscr{R}(\widehat{p}, p) := \int_{[-1,1]} \mathbb{E}[\widehat{p}(x_0) - p(x_0)]$, the estimator achieves the minimax risk $n^{-2s/(2s+1)}$ over $\mathscr{H}_s(C, L)$ up to a $\sqrt{\log n}$ factor.

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See paper for a Gaussian mixture example.



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 - Cryptographic hardness results for learning a sampler remain elusive.

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Our reduction framework yields a *general* blueprint:

To prove crypto hardness for learning the scores of a family $\mathcal{P}:$

- 1. Check that ${\mathscr P}$ satisfies the conditions of our reduction.
- 2. Prove that PAC density estimation over ${\mathscr P}$ is crypto hard.

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Recently, [Song '24] proved crypto hardness of learning score functions for Gaussian mixtures via a tailored argument.

Our reduction framework yields a *general* blueprint:

To prove crypto hardness for learning the scores of a family $\mathcal{P} {:}$

- 1. Check that \mathcal{P} satisfies the conditions of our reduction.
- 2. Prove that PAC density estimation over $\mathcal P$ is crypto hard.

Learning the scores is as <u>hard</u> as PAC density estimation.

Application to Gaussian mixtures

Corollary [CKMM '25]: For any $\varepsilon > 0$, it is cryptographically hard to learn the score functions of mixtures of Gaussians with up to d^{ε} components.

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Reduction chain for experts:

- score estimation ← PAC density estimation (our framework)
- ← CLWE (following [Bruna, Regev, Song, Tang '21])
- ← LWE (following [Gupte, Vafa, Vaikuntanathan '22])
- ← lattice problems [Regev '09]
- ← post-quantum cryptography.

Summary

